

# On Clustering Histograms with $k$ -Means by Using Mixed $\alpha$ -Divergences

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# Clustering histograms

- ▶ Information Retrieval systems (IRs) based on **bag-of-words** paradigm (bag-of-textons, bag-of-features, bag-of-X)
- ▶ The rôle of distances:
  - ▶ Initially, create a dictionary of “words” by quantizing using  $k$ -means clustering (depends on the underlying distance)
  - ▶ At query time, find “closest” (histogram) document by querying with the histogram query
- ▶ Notation: Positive arrays  $h$  (counting histogram) versus frequency histograms  $\tilde{h}$  (normalized counting)  $d$  bins

For IRs, prefer **symmetric distances** (not necessarily metrics) like the Jeffreys divergence or the Jensen-Shannon divergence (unified by a one parameterized family of divergences in [11])

## Ali-Silvey-Csiszár $f$ -divergences

An important class of divergences:  $f$ -divergences [10, 1, 7] defined for a convex generator  $f$  (with  $f(1) = f'(1) = 0$  and  $f''(1) = 1$ ):

$$I_f(p : q) \doteq \sum_{i=1}^d q^i f\left(\frac{p^i}{q^i}\right)$$

Those divergences preserve **information monotonicity** [3] under any arbitrary transition probability (Markov morphisms).  
 $f$ -divergences can be extended to positive arrays [3].

# Mixed divergences

Defined on three parameters:

$$M_\lambda(p : q : r) \doteq \lambda D(p : q) + (1 - \lambda) D(q : r)$$

for  $\lambda \in [0, 1]$ .

Mixed divergences include:

- ▶ the **sided divergences** for  $\lambda \in \{0, 1\}$ ,
- ▶ the **symmetrized** (arithmetic mean) divergence for  $\lambda = \frac{1}{2}$ .

## Mixed divergence-based $k$ -means clustering

$k$  distinct seeds from the dataset with  $l_i = r_i$ .

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**Input:** Weighted histogram set  $\mathcal{H}$ , divergence  $D(\cdot, \cdot)$ , integer  $k > 0$ , real  $\lambda \in [0, 1]$ ;

Initialize left-sided/right-sided seeds  $\mathcal{C} = \{(l_i, r_i)\}_{i=1}^k$ ;

**repeat**

  //Assignment

**for**  $i = 1, 2, \dots, k$  **do**

$\mathcal{C}_i \leftarrow \{h \in \mathcal{H} : i = \arg \min_j M_\lambda(l_j : h : r_j)\}$ ;

  // Dual-sided centroid relocation

**for**  $i = 1, 2, \dots, k$  **do**

$r_i \leftarrow \arg \min_x D(\mathcal{C}_i : x) = \sum_{h \in \mathcal{C}_i} w_j D(h : x)$ ;

$l_i \leftarrow \arg \min_x D(x : \mathcal{C}_i) = \sum_{h \in \mathcal{C}_i} w_j D(x : h)$ ;

**until** convergence;

**Output:** Partition of  $\mathcal{H}$  into  $k$  clusters following  $\mathcal{C}$ ;

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→ different from the  $k$ -means clustering with respect to the symmetrized divergences

## $\alpha$ -divergences

For  $\alpha \in \mathbb{R} \neq \pm 1$ , define  $\alpha$ -divergences [6] on positive arrays [18] :

$$D_\alpha(p : q) \doteq \sum_{i=1}^d \frac{4}{1 - \alpha^2} \left( \frac{1 - \alpha}{2} p^i + \frac{1 + \alpha}{2} q^i - (p^i)^{\frac{1 - \alpha}{2}} (q^i)^{\frac{1 + \alpha}{2}} \right)$$

with  $D_\alpha(p : q) = D_{-\alpha}(q : p)$  and in the limit cases  $D_{-1}(p : q) = \text{KL}(p : q)$  and  $D_1(p : q) = \text{KL}(q : p)$ , where KL is the extended Kullback–Leibler divergence:

$$\text{KL}(p : q) \doteq \sum_{i=1}^d p^i \log \frac{p^i}{q^i} + q^i - p^i.$$

## $\alpha$ -divergences belong to $f$ -divergences

The  $\alpha$ -divergences belong to the class of Csiszár  $f$ -divergences with the following generator:

$$f(t) = \begin{cases} \frac{4}{1-\alpha^2} (1 - t^{(1+\alpha)/2}), & \text{if } \alpha \neq \pm 1, \\ t \ln t, & \text{if } \alpha = 1, \\ -\ln t, & \text{if } \alpha = -1 \end{cases}$$

The Pearson and Neyman  $\chi^2$  distances are obtained for  $\alpha = -3$  and  $\alpha = 3$ :

$$D_3(\tilde{p} : \tilde{q}) = \frac{1}{2} \sum_i \frac{(\tilde{q}^i - \tilde{p}^i)^2}{\tilde{p}^i},$$

$$D_{-3}(\tilde{p} : \tilde{q}) = \frac{1}{2} \sum_i \frac{(\tilde{q}^i - \tilde{p}^i)^2}{\tilde{q}^i}.$$

## Squared Hellinger symmetric distance is a $\alpha = 0$ -divergence

Divergence  $D_0$  is the squared Hellinger symmetric distance (scaled by 4) extended to positive arrays:

$$D_0(p : q) = 2 \int \left( \sqrt{p(x)} - \sqrt{q(x)} \right)^2 dx = 4H^2(p, q),$$

with the Hellinger distance:

$$H(p, q) = \sqrt{\frac{1}{2} \int \left( \sqrt{p(x)} - \sqrt{q(x)} \right)^2 dx}$$



## Mixed $\alpha$ -divergences

- ▶ Mixed  $\alpha$ -divergence between a histogram  $x$  to **two** histograms  $p$  and  $q$ :

$$\begin{aligned}M_{\lambda,\alpha}(p : x : q) &= \lambda D_{\alpha}(p : x) + (1 - \lambda) D_{\alpha}(x : q), \\ &= \lambda D_{-\alpha}(x : p) + (1 - \lambda) D_{-\alpha}(q : x), \\ &= M_{1-\lambda,-\alpha}(q : x : p),\end{aligned}$$

- ▶  $\alpha$ -Jeffreys symmetrized divergence is obtained for  $\lambda = \frac{1}{2}$ :

$$S_{\alpha}(p, q) = M_{\frac{1}{2},\alpha}(q : p : q) = M_{\frac{1}{2},\alpha}(p : q : p)$$

- ▶ skew symmetrized  $\alpha$ -divergence is defined by:

$$S_{\lambda,\alpha}(p : q) = \lambda D_{\alpha}(p : q) + (1 - \lambda) D_{\alpha}(q : p)$$

# Coupled $k$ -Means++ $\alpha$ -Seeding

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**Algorithm 1:** Mixed  $\alpha$ -seeding; MAS( $\mathcal{H}$ ,  $k$ ,  $\lambda$ ,  $\alpha$ )

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**Input:** Weighted histogram set  $\mathcal{H}$ , integer  $k \geq 1$ , real  $\lambda \in [0, 1]$ ,  
real  $\alpha \in \mathbb{R}$ ;

Let  $\mathcal{C} \leftarrow h_j$  with uniform probability ;

**for**  $i = 2, 3, \dots, k$  **do**

    Pick at random histogram  $h \in \mathcal{H}$  with probability:

$$\pi_{\mathcal{H}}(h) \doteq \frac{w_h M_{\lambda, \alpha}(c_h : h : c_h)}{\sum_{y \in \mathcal{H}} w_y M_{\lambda, \alpha}(c_y : y : c_y)}, \quad (1)$$

    //where  $(c_h, c_h) \doteq \arg \min_{(z, z) \in \mathcal{C}} M_{\lambda, \alpha}(z : h : z)$ ;

$\mathcal{C} \leftarrow \mathcal{C} \cup \{(h, h)\}$ ;

**Output:** Set of initial cluster centers  $\mathcal{C}$ ;

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## A guaranteed probabilistic initialization

Let  $C_{\lambda,\alpha}$  denote for short the cost function related to the clustering type chosen (left-, right-, skew Jeffreys or mixed) in MAS and  $C_{\lambda,\alpha}^{opt}$  denote the optimal related clustering in  $k$  clusters, for  $\lambda \in [0, 1]$  and  $\alpha \in (-1, 1)$ . Then, on average, with respect to distribution (1), the initial clustering of MAS satisfies:

$$E_{\pi}[C_{\lambda,\alpha}] \leq 4 \begin{cases} f(\lambda)g(k)h^2(\alpha)C_{\lambda,\alpha}^{opt} & \text{if } \lambda \in (0, 1) \\ g(k)z(\alpha)h^4(\alpha)C_{\lambda,\alpha}^{opt} & \text{otherwise} \end{cases} .$$

Here,  $f(\lambda) = \max \left\{ \frac{1-\lambda}{\lambda}, \frac{\lambda}{1-\lambda} \right\}$ ,  $g(k) = 2(2 + \log k)$ ,  $z(\alpha) = \left( \frac{1+|\alpha|}{1-|\alpha|} \right)^{\frac{8|\alpha|^2}{(1-|\alpha|)^2}}$ ,  $h(\alpha) = \max_i p_i^{|\alpha|} / \min_i p_i^{|\alpha|}$ ; the min is defined on strictly positive coordinates, and  $\pi$  denotes the picking distribution.

## Mixed $\alpha$ -hard clustering: MAhC( $\mathcal{H}$ , $k$ , $\lambda$ , $\alpha$ )

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**Input:** Weighted histogram set  $\mathcal{H}$ , integer  $k > 0$ , real  $\lambda \in [0, 1]$ ,  
real  $\alpha \in \mathbb{R}$ ;

Let  $\mathcal{C} = \{(l_i, r_i)\}_{i=1}^k \leftarrow \text{MAS}(\mathcal{H}, k, \lambda, \alpha)$ ;

**repeat**

  //Assignment

**for**  $i = 1, 2, \dots, k$  **do**

$\mathcal{A}_i \leftarrow \{h \in \mathcal{H} : i = \arg \min_j M_{\lambda, \alpha}(l_j : h : r_j)\}$ ;

  // Centroid relocation

**for**  $i = 1, 2, \dots, k$  **do**

$r_i \leftarrow \left( \sum_{h \in \mathcal{A}_i} w_i h^{\frac{1-\alpha}{2}} \right)^{\frac{2}{1-\alpha}}$ ;

$l_i \leftarrow \left( \sum_{h \in \mathcal{A}_i} w_i h^{\frac{1+\alpha}{2}} \right)^{\frac{2}{1+\alpha}}$ ;

**until** convergence;

**Output:** Partition of  $\mathcal{H}$  in  $k$  clusters following  $\mathcal{C}$ ;

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## Sided Positive $\alpha$ -Centroids [14]

The left-sided  $l_\alpha$  and right-sided  $r_\alpha$  positive weighted  $\alpha$ -centroid coordinates of a set of  $n$  positive histograms  $h_1, \dots, h_n$  are weighted  $\alpha$ -means:

$$r_\alpha^i = f_\alpha^{-1} \left( \sum_{j=1}^n w_j f_\alpha(h_j^i) \right), l_\alpha^i = r_{-\alpha}^i$$

$$\text{with } f_\alpha(x) = \begin{cases} x^{\frac{1-\alpha}{2}} & \alpha \neq \pm 1, \\ \log x & \alpha = 1. \end{cases}$$

## Sided Frequency $\alpha$ -Centroids [2]

Theorem (Amari, 2007)

*The coordinates of the sided frequency  $\alpha$ -centroids of a set of  $n$  weighted frequency histograms are the normalised weighted  $\alpha$ -means.*

# Positive and Frequency $\alpha$ -centroids

Summary:

- ▶  $r_{\alpha}^i = \begin{cases} (\sum_{j=1}^n w_j (h_j^i)^{\frac{1-\alpha}{2}})^{\frac{2}{1-\alpha}} & \alpha \neq 1 \\ r_1^i = \prod_{j=1}^n (h_j^i)^{w_j} & \alpha = 1 \end{cases}$
- ▶  $l_{\alpha}^i = r_{-\alpha}^i = \begin{cases} (\sum_{j=1}^n w_j (h_j^i)^{\frac{1+\alpha}{2}})^{\frac{2}{1+\alpha}} & \alpha \neq -1 \\ l_{-1}^i = \prod_{j=1}^n (h_j^i)^{w_j} & \alpha = -1 \end{cases}$
- ▶  $\tilde{r}_{\alpha}^i = \frac{r_{\alpha}^i}{w(\tilde{r}_{\alpha})}$
- ▶  $\tilde{l}_{\alpha}^i = \tilde{r}_{-\alpha}^i = \frac{r_{-\alpha}^i}{w(\tilde{r}_{-\alpha})}$

# Mixed $\alpha$ -Centroids

Two centroids minimizer of:

$$\sum_j w_j M_{\lambda, \alpha}(l : h_j : r)$$

Generalizing mixed Bregman divergences [16]:

## Theorem

*The two mixed  $\alpha$ -centroids are the left-sided and right-sided  $\alpha$ -centroids.*



## Symmetrized Jeffreys-Type $\alpha$ -Centroids

$$\begin{aligned} S_\alpha(p, q) &= \frac{1}{2} (D_\alpha(p : q) + D_\alpha(q : p)) = S_{-\alpha}(p, q), \\ &= M_{\frac{1}{2}}(p : q : p), \end{aligned}$$

For  $\alpha = \pm 1$ , we get half of Jeffreys divergence:

$$S_{\pm 1}(p, q) = \frac{1}{2} \sum_{i=1}^d (p^i - q^i) \log \frac{p^i}{q^i}$$

## Jeffreys $\alpha$ -divergence and Heinz means

When  $p$  and  $q$  are frequency histograms, we have for  $\alpha \neq \pm 1$ :

$$J_\alpha(\tilde{p} : \tilde{q}) = \frac{8}{1 - \alpha^2} \left( 1 + \sum_{i=1}^d H_{\frac{1-\alpha}{2}}(\tilde{p}^i, \tilde{q}^i) \right)$$

where  $H_{\frac{1-\alpha}{2}}(a, b)$  a symmetric Heinz mean [8, 5]:

$$H_\beta(a, b) = \frac{a^\beta b^{1-\beta} + a^{1-\beta} b^\beta}{2}$$

Heinz means interpolate the arithmetic and geometric means and satisfies the inequality:

$$\sqrt{ab} = H_{\frac{1}{2}}(a, b) \leq H_\alpha(a, b) \leq H_0(a, b) = \frac{a+b}{2}.$$

## Jeffreys divergence in the limit case

For  $\alpha = \pm 1$ ,  $S_\alpha(p, q)$  tends to the Jeffreys divergence:

$$J(p, q) = \text{KL}(p, q) + \text{KL}(q, p) = \sum_{i=1}^d (p^i - q^i)(\log p^i - \log q^i)$$

The Jeffreys divergence writes mathematically the same for frequency histograms:

$$J(\tilde{p}, \tilde{q}) = \text{KL}(\tilde{p}, \tilde{q}) + \text{KL}(\tilde{q}, \tilde{p}) = \sum_{i=1}^d (\tilde{p}^i - \tilde{q}^i)(\log \tilde{p}^i - \log \tilde{q}^i)$$

## Analytic formula for the positive Jeffreys centroid [12]

### Theorem (Jeffreys positive centroid [12])

*The Jeffreys positive centroid  $c = (c^1, \dots, c^d)$  of a set  $\{h_1, \dots, h_n\}$  of  $n$  weighted positive histograms with  $d$  bins can be calculated component-wise exactly using the Lambert  $W$  analytic function:*

$$c^i = \frac{a^i}{W\left(\frac{a^i}{g^i} e\right)}$$

*where  $a^i = \sum_{j=1}^n \pi_j h_j^i$  denotes the coordinate-wise arithmetic weighted means and  $g^i = \prod_{j=1}^n (h_j^i)^{\pi_j}$  the coordinate-wise geometric weighted means.*

The Lambert analytic function  $W$  [4] (positive branch) is defined by  $W(x)e^{W(x)} = x$  for  $x \geq 0$ .

## Jeffreys frequency centroid [12]

### Theorem (Jeffreys frequency centroid [12])

Let  $\tilde{c}$  denote the Jeffreys frequency centroid and  $\tilde{c}' = \frac{c}{w_c}$  the normalised Jeffreys positive centroid. Then, the approximation factor  $\alpha_{\tilde{c}'} = \frac{S_1(\tilde{c}', \tilde{\mathcal{H}})}{S_1(\tilde{c}, \tilde{\mathcal{H}})}$  is such that  $1 \leq \alpha_{\tilde{c}'} \leq \frac{1}{w_c}$  (with  $w_c \leq 1$ ).

better upper bounds in [12].

## Reducing a $n$ -size problem to a 2-size problem

Generalize [17] (symmetrized Kullback–Leibler divergence) and [15] (symmetrized Bregman divergence)

### Lemma (Reduction property)

*The symmetrized  $J_\alpha$ -centroid of a set of  $n$  weighted histograms amount to computing the symmetrized  $\alpha$ -centroid for the weighted  $\alpha$ -mean and  $-\alpha$ -mean:*

$$\min_x J_\alpha(x, \mathcal{H}) = \min_x (D_\alpha(x : r_\alpha) + D_\alpha(l_\alpha : x)).$$

## Frequency symmetrized $\alpha$ -centroid

Minimizer of  $\min_{\tilde{x} \in \Delta_d} \sum_j w_j S_\alpha(\tilde{x}, \tilde{h}_j)$

Instead of seeking for  $\tilde{x}$  in the probability simplex, we can optimize on the unconstrained domain  $\mathbb{R}^{d-1}$  by using the natural parameter reparameterization [13] of multinomials.

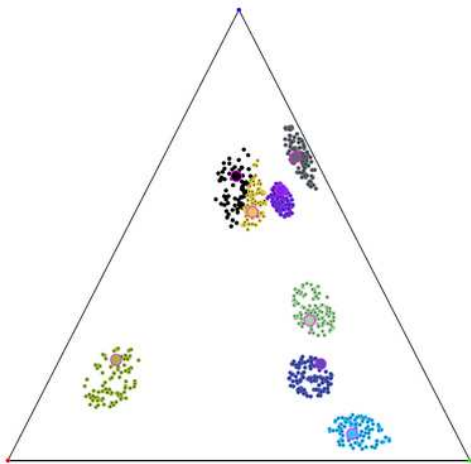
### Lemma

*The  $\alpha$ -divergence for distributions belonging to the same exponential families amounts to computing a divergence on the corresponding natural parameters:*

$$A_\alpha(p : q) = \frac{4}{1 - \alpha^2} \left( 1 - e^{-J_F^{\left(\frac{1-\alpha}{2}\right)}(\theta_p : \theta_q)} \right),$$

*where  $J_F^\beta(\theta_1 : \theta_2) = \beta F(\theta_1) + (1 - \beta)F(\theta_2) - F(\beta\theta_1 + (1 - \beta)\theta_2)$  is a skewed Jensen divergence defined for the log-normaliser  $F$  of the family.*

## Implementation (in `processing.org`)



Snapshot of the  $\alpha$ -clustering software. Here,  $n = 800$  frequency histograms of three bins with  $k = 8$ , and  $\alpha = 0.7$  and  $\lambda = \frac{1}{2}$ .



# Soft Mixed $\alpha$ -Clustering

Learn both  $\alpha$  and  $\lambda$  ( $\alpha$ -EM [9])

**Input:** Histogram set  $\mathcal{H}$  with  $|\mathcal{H}| = m$ , integer  $k > 0$ , real

$\lambda \leftarrow \lambda_{\text{init}} \in [0, 1]$ , real  $\alpha \in \mathbb{R}$ ;

Let  $\mathcal{C} = \{(l_j, r_i)\}_{i=1}^k \leftarrow \text{MAS}(\mathcal{H}, k, \lambda, \alpha)$ ;

**repeat**

//Expectation

**for**  $i = 1, 2, \dots, m$  **do**

**for**  $j = 1, 2, \dots, k$  **do**

$$\left[ \left[ \rho(j|h_i) = \frac{\pi_j \exp(-M_{\lambda, \alpha}(l_j : h_i : r_j))}{\sum_{j'} \pi_{j'} \exp(-M_{\lambda, \alpha}(l_{j'} : h_i : r_{j'}))} \right] \right];$$

//Maximization

**for**  $j = 1, 2, \dots, k$  **do**

$$\left[ \left[ \pi_j \leftarrow \frac{1}{m} \sum_i \rho(j|h_i); \right. \right.$$

$$\left. \left. l_i \leftarrow \left( \frac{1}{\sum_i \rho(j|h_i)} \sum_i \rho(j|h_i) h_i^{\frac{1+\alpha}{2}} \right)^{\frac{2}{1+\alpha}}; \right. \right.$$

$$\left. \left. r_i \leftarrow \left( \frac{1}{\sum_i \rho(j|h_i)} \sum_i \rho(j|h_i) h_i^{\frac{1-\alpha}{2}} \right)^{\frac{2}{1-\alpha}}; \right. \right.$$

//Alpha - Lambda

$$\alpha \leftarrow \alpha - \eta_1 \sum_{j=1}^k \sum_{i=1}^m \rho(j|h_i) \frac{\partial}{\partial \alpha} M_{\lambda, \alpha}(l_j : h_i : r_j);$$

**if**  $\lambda_{\text{init}} \neq 0, 1$  **then**

$$\left[ \left[ \lambda \leftarrow \lambda - \eta_2 \left( \sum_{j=1}^k \sum_{i=1}^m \rho(j|h_i) D_{\alpha}(l_j : h_i) - \right. \right. \right.$$

$$\left. \left. \sum_{j=1}^k \sum_{i=1}^m \rho(j|h_i) D_{\alpha}(h_i : r_j) \right); \right.$$

    //for some small  $\eta_1, \eta_2$ ; ensure that  $\lambda \in [0, 1]$ .

**until** convergence;

**Output:** Soft clustering of  $\mathcal{H}$  according to  $k$  densities  $\rho(j|\cdot)$

following  $\mathcal{C}$ ;

# Summary

1. Mixed divergences, mixed divergence  $k$ -means++ seeding, coupled  $k$ -means seeding
2. Sided left or right  $\alpha$ -centroid  $k$ -means
3. Coupled  $k$ -means with respect to mixed  $\alpha$ -divergences relying on dual  $\alpha$ -centroids
4. Symmetrized Jeffreys-type  $\alpha$ -centroid (variational)  $k$ -means,

All technical proofs and details in:  
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